## **Chebyshev's Theorem**

The probability that any random variable X falls within k standard deviations of the mean is at least

$$(1 - \frac{1}{k^2})$$
. That is

$$\Pr(\mu - k\sigma < x < \mu + k\sigma) \ge 1 - \frac{1}{k^2}$$

## **Proof**

$$\sigma^{2} = E \Big[ (x - \mu)^{2} \Big]$$

$$= \int_{-\infty}^{\infty} (x - \mu)^{2} f(x) dx$$

$$= \int_{-\infty}^{\mu - k\sigma} (x - \mu)^{2} f(x) dx + \int_{\mu - k\sigma}^{\mu + k\sigma} (x - \mu)^{2} f(x) dx + \int_{\mu + k\sigma}^{\infty} (x - \mu)^{2} f(x) dx$$

$$\geq \int_{-\infty}^{\mu - k\sigma} (x - \mu)^{2} f(x) dx + \int_{\mu + k\sigma}^{\infty} (x - \mu)^{2} f(x) dx$$

Since the 2<sup>nd</sup> of the three integrals is non-negative.

Now since  $|x-\mu| \ge k\sigma$  wherever  $x \ge \mu + k\sigma$  or  $x \le \mu - k\sigma$ , we have  $(x-\mu)^2 \ge k^2\sigma^2$  in both remaining integrals. It follows that

$$\sigma^2 \ge \int_{-\infty}^{\mu - k\sigma} k^2 \sigma^2 f(x) dx + \int_{\mu + k\sigma}^{\infty} k^2 \sigma^2 f(x) dx$$

and that

$$\frac{1}{k^2} \ge \int_{-\infty}^{\mu - k\sigma} f(x) dx + \int_{\mu + k\sigma}^{\infty} f(x) dx$$

$$\therefore 1 - \frac{1}{k^2} \le 1 - \int_{-\infty}^{\mu - k\sigma} f(x) dx - \int_{\mu + k\sigma}^{\infty} f(x) dx$$

$$1 - \frac{1}{k^2} \le \int_{\mu - k\sigma}^{\mu + k\sigma} f(x) dx = \Pr(\mu - k\sigma < x < \mu + k\sigma)$$